

## Advanced insurance.

Consider a risky situation where the loss (indicated with positive values  $x \geq 0$ ) is a continuous r.v.  $L$  with density  $f(x)$ . We assume that  $E(L) > 0$ , which means that  $L$  is not always 0.

We define  $I(x)$  what the insurance company pays to the customer in case of loss  $x$ . Clearly  $0 \leq I(x) \leq x$  since no customer would accept negative reimbursements (therefore having to pay money to the insurance company instead of receiving) and no insurance company reimburse more than the effective loss.

$I(x)$  is

- full coverage:  $I(x) = x$
- proportional:  $I(x) = kx$ , where  $k \in (0;1)$
- limited coverage:  $I(x) = \begin{cases} x \leq m & x \\ x > m & m \end{cases}$
- stop loss:  $I(x) = \begin{cases} x \leq d & 0 \\ x > d & x - d \end{cases}$ , where  $d$  is the deductible.

We define the **expected claim** as the expected value of the money received by the customer. It is sometimes called **expected loss** since for the insurance company it represents a loss. Its definition is  $E(I(x)) = \int_0^{+\infty} I(s)f(s)ds$ . On the expected claim we have the restriction that it must be positive or zero since both functions are positive or zero.

Moreover,  $\int_0^{+\infty} I(s)f(s)ds \leq \int_0^{+\infty} sf(s)ds = E(L)$ .

For the stop-loss insurance the expected claim becomes

$$E(I(x)) = \int_0^d 0 f(s)ds + \int_d^{+\infty} (s-d)f(s)ds = \int_d^{+\infty} (s-d)f(s)ds$$

Q.2

$$R(r) = E(I_r(L)) \quad \text{DEFINITION}$$

$$\begin{aligned} R(r) &= \int_r^{+\infty} (s-r) f(s) ds = \int_0^{+\infty} (s-r) f(s) ds - \int_0^r (s-r) f(s) ds = \\ &= \int_0^{+\infty} s f(s) ds - \int_0^r s f(s) ds - r \int_0^{+\infty} f(s) ds + r \int_0^r f(s) ds = \\ &= \left[ E(L) - \int_0^r s f(s) ds \right] - \left[ r - r \int_0^r f(s) ds \right] = R_1(r) - R_2(r) \end{aligned}$$

$$R'(r) = -r f(r) - 1 + \int_0^r f(s) ds + r \cdot f(r) = \int_0^r f(s) ds - 1$$

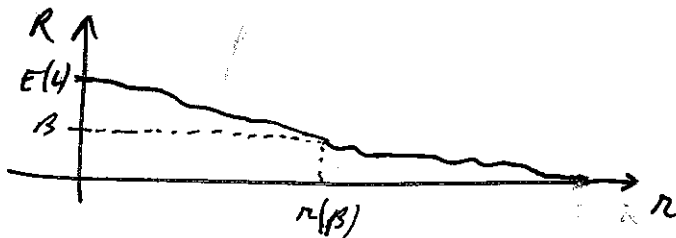
SINCE  $\int_0^r f(s) < 1 \Rightarrow R'(r) < 0 \Rightarrow R$  STRICTLY DECREASING

$$R(0) = \int_0^{+\infty} (s-0) f(s) ds = E(L) > 0 \quad \text{SINCE THERE IS A LOSS IN THIS LOTTERY}$$

$$\lim_{r \rightarrow +\infty} R(r) = \lim_{r \rightarrow +\infty} \int_r^{+\infty} (s-r) f(s) ds = \underline{\underline{0}}$$

WE DO NOT SHOW THE PROOF

$R$  IS DECREASING, STARTS FROM  $E(L)$  AND GOES TO  $0$



THEREFORE FOR ANY  $\beta \in (0, E(L))$  THERE IS A CORRESPONDING  $r$ , CALLED  $r(\beta)$ , WHICH

$$R(r(\beta)) = \beta$$

AND THIS  $r(\beta)$  IS UNIQUE.

9.3)

EXAMPLE:

$$L \sim \lambda e^{-\lambda x}$$

STOP-LOSS INSURANCE

FIND  $r(\beta)$

NOTE THAT  $E(L) = \frac{1}{\lambda}$

$$R(r) = E(I_r(L)) = \int_r^{+\infty} (s-r) \lambda e^{-\lambda s} ds = \text{BY PARTS}$$

$$= \left[ (s-r) \left( -\frac{1}{\lambda} \right) \lambda e^{-\lambda s} \right]_r^{+\infty} - \int_r^{+\infty} \left( -\frac{1}{\lambda} \right) \lambda e^{-\lambda s} ds =$$

$$= 0 + \frac{1}{\lambda} (r-r) \lambda e^{-\lambda r} + \left[ -\frac{e^{-\lambda s}}{\lambda} \right]_r^{+\infty} =$$

$$= 0 + 0 - \frac{1}{\lambda} \left[ e^{-\lambda s} \right]_r^{+\infty} = 0 + \frac{1}{\lambda} e^{-\lambda r}$$

$$R(r) = \frac{1}{\lambda} e^{-\lambda r}$$

$$R(r(\beta)) = \beta$$

$$\frac{1}{\lambda} e^{-\lambda r(\beta)} = \beta$$

$$e^{-\lambda r(\beta)} = \beta \lambda$$

$$-\lambda r(\beta) = \ln \beta \lambda$$

$$r(\beta) = -\frac{1}{\lambda} \ln \beta \lambda$$

FOR AN EXPECTED CLAIM OF  $\beta$  THE DEDUCTIBLE MUST BE  $-\frac{1}{\lambda} \ln \beta \lambda$

NOTE THAT  $\beta \leq \frac{1}{\lambda}$  SINCE  $E(L) = \frac{1}{\lambda}$  AND  $R$  MUST BE  $\leq E(L)$  NOT

$$\ln \beta \lambda \leq 0 \Rightarrow r(\beta) \geq 0$$

9.4

## EXERCISE

STOP-LOSS INSURANCE WITH DEDUCTIBLE  $d$

LOSS IS UNIFORMLY DISTRIBUTED OVER  $[0, 3d]$

WHAT IS EXPECTED CLAIM?

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$$\text{EXPECTED CLAIM} = E(I_d(L)) = \int_0^{+\infty} I_d(s) f(s) ds = \int_0^d 0 f(s) ds + \int_d^{3d} (s-d) f(s) ds$$

$$f(s) = \begin{cases} \frac{1}{3d} & s \in [0, 3d] \\ 0 & s \notin [0, 3d] \end{cases}$$

$$I_d(s) = \begin{cases} 0 & s \leq d \\ s-d & s > d \end{cases}$$

$$E(I_d(L)) = \int_d^{3d} (s-d) \cdot \frac{1}{3d} ds = \left[ \frac{(s-d)^2}{6d} \right]_d^{3d} = \frac{(3d-d)^2}{6d} - \frac{(d-d)^2}{6d} = \frac{4d^2}{6d} = \frac{2}{3}d$$

□

9.5

EXERCISE

HW STOP-LOSS INSURANCE

LOSS is  $U(0; 5)$   $E(I_d(U)) = 1$ FIND  $d$ 

$$1 = E(I_d(U)) = \int_0^d 0 \cdot f(s) ds + \int_d^5 (s-d) f(s) ds =$$

$$f(s) = \begin{cases} \frac{1}{5} & s \in [0, 5] \\ 0 & s \notin [0, 5] \end{cases}$$

$$I_d(s) = \begin{cases} 0 & s \leq d \\ s-d & s > d \end{cases}$$

$$= \int_d^5 (s-d) \frac{1}{5} ds = \left[ \frac{(s-d)^2}{10} \right]_d^5 = \frac{(5-d)^2}{10} - \frac{(d-d)^2}{10} = \frac{(5-d)^2}{10}$$

$$\frac{(5-d)^2}{10} = 1 \quad (5-d)^2 = 10 \quad 5-d = \pm\sqrt{10}$$

$$d = 5 \mp \sqrt{10} = \begin{cases} 5 + \sqrt{10} & \rightarrow \text{THIS VALUE DOES NOT MAKE SENSE} \\ & \text{BECAUSE } d \text{ MAY NOT BE LARGER THAN} \\ & \text{THE MAXIMUM LOSS!} \\ 5 - \sqrt{10} \end{cases}$$

$$d = 5 - \sqrt{10}$$

D

9.6

# OPTIMAL INSURANCE THEOREM

DECISION MAKER WITH

- 1) RISK AVERSE WITH  $U$  TWICE-DIFFERENTIABLE
- 2)  $L$  IS A LOSS WITH  $f(x)$  DENSITY FUNCTION  $x \in [0, +\infty)$
- 3) READY TO SPEND  $P$  ON AN INSURANCE
- 4) AMONG ALL CONTRACTS  $I(x)$  WITH  $0 \leq I(x) \leq x$   
WITH  $E(I(L)) = \beta$  EXPECTED CLAIM

HIS EXPECTED UTILITY IS MAXIMIZED CHOOSING A

STOP-LOSS CONTRACT WITH

DEDUCTIBLE  $n$

SUCH THAT

$$\beta = \int_n^{+\infty} (s-n) f(s) ds$$

NOTES:

\*  $n$  DOES NOT DEPEND ON  $U$  PROVIDED THAT  $U'' \leq 0$

\* WHAT DOES MAXIMIZING EXPECTED UTILITY MEAN?  
IN THIS CASE IT MEANS MAXIMIZING THE AMOUNT  
OF MONEY RECEIVED FROM THE INSURANCE COMPANY

$$E(U(I(L))) = \int_0^{+\infty} U(I(s) + w_0) \cdot f(s) ds$$

WITH THE BOUND THAT  $\int_0^{+\infty} I(s) \cdot f(s) ds = \beta$

WHICH, FOR A STOP LOSS, MEANS  $\int_n^{+\infty} (s-n) f(s) ds = \beta$

NOTE:  $\beta$  IS ALSO CALLED EXPECTED LOSS

9.7 IMAGINE A RISKY SITUATION WITH  $R(0;7)$

CONSIDER ALL INSURANCES WITH EXPECTED CLAIM EQUAL TO 0,2. SUPPOSE  $U(x)$  TWICE DIFFERENTIABLE  $U'' < 0$ . SUPPOSE ALL INSURANCES REQUIRE THE SAME PREMIUM. FIND THE BEST

USING OPTIMAL INSURANCE THEORY, SINCE  $U$  IS TWICE DIFF, AMONG ALL CONTRACTS WITH  $0 \leq I(x) \leq x$  WITH  $E(I(x)) = 0,2$ , THE EXPECTED UTILITY IS MAXIMIZED CHOOSING

$$0,2 = \int_d^{+\infty} (s-d) f(s) ds = \int_d^7 (s-d) \cdot \frac{1}{7} ds = \left[ \frac{(s-d)^2}{14} \right]_d^7 = \frac{(7-d)^2}{14} - 0 = \frac{49-14d+d^2}{14}$$

$$2,8 = 49 - 14d + d^2$$

$$d^2 - 14d + 46,2 = 0$$

$$d_{1,2} = \frac{7 \pm \sqrt{49 - 46,2}}{1} = 7 \pm \sqrt{2,8} = \begin{cases} 8,67 \rightarrow \text{NOT POSSIBLE} \\ 5,33 \end{cases} \begin{matrix} \text{SINCE LOSS} \\ \text{IS } R(0;7) \end{matrix}$$

$d = 5,33$  IS THE BEST FEASIBLE

9.8

U TWICE-DIFFERENTIABLE

 $U'' < 0$ 

HW

$$L \sim \lambda e^{-\lambda x} \quad \lambda > 0$$

EXPECTED CLAIM IS  $\frac{1}{2\lambda}$

WHAT IS THE BEST POSSIBLE INSURANCE CONTRACT

ACCORDING TO THE OPTIMAL INSURANCE THEOREM

IF AGENT IS RISK-AVERSE THE OPTIMAL CONTRACT IS A STOP-LOSS WITH DEDUCTIBLE  $d$

$$\frac{1}{2\lambda} = \int_d^{+\infty} (s-d) \lambda e^{-\lambda s} ds$$

$$\int_d^{+\infty} s \lambda e^{-\lambda s} ds - d \int_d^{+\infty} \lambda e^{-\lambda s} ds = \left[ s \cdot (-e^{-\lambda s}) \right]_d^{+\infty} - \int_d^{+\infty} 1 \cdot (-e^{-\lambda s}) ds - d \left[ -e^{-\lambda s} \right]_d^{+\infty} =$$

$$= 0 + \cancel{d e^{-\lambda d}} + \left[ \frac{-e^{-\lambda s}}{\lambda} \right]_d^{+\infty} + d \left( 0 - \cancel{e^{-\lambda d}} \right) = + \frac{e^{-\lambda d}}{\lambda}$$

$$\frac{1}{2\lambda} = \frac{e^{-\lambda d}}{\lambda}$$

$$\ln \frac{1}{2} = -\lambda d$$

$$d = -\frac{1}{\lambda} \ln \frac{1}{2} = \frac{1}{\lambda} \ln 2$$

□

9.9  $I(x) = \sqrt{x}$  DOES THIS CORRESPOND TO  
A FEASIBLE INSURANCE CONTRACT?

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ON  $I(x)$  WE HAVE THE RESTRICTIONS

$$0 \leq I(x) \leq x$$

$0 \leq I(x)$  IS TRUE FOR  $\sqrt{x}$

BUT  $\sqrt{x} \leq x$  IS NOT TRUE (FOR EXAMPLE  
 $x = 0,5$  DOES NOT WORK)

□